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AN ADAPTATION OF A MARKOV
CHAIN MODEL FOR ANTISUBMARINE WARFARE
CARRIER AIRCRAFT

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AN ADAPTATION OF A
MARKOV CHAIN MODEL FOR
ANTISUBMARINE WARFARE CARRIER AIRCRAFT

by

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ABSTRACT

It is the purpose of this paper to develop a useful mathematical model of ASW aircraft availability. The increasing emphasis of systems studies dictates the use of accurate and representative models of the ASW systems. At present, many studies are using essentially the same models developed during World War II. This paper is an attempt to make use of advanced theory in a more powerful and flexible model and to make the use of the model practical and verifiable.

The writer adapted the time homogeneous bivariate model as developed by F. C. Collins. This is a discrete time Markov process with a stochastic matrix of transition probabilities wherein the maintenance process is modeled as a pulsed input multiple server queue.

The model was programmed in FORTRAN 63 on the CDC 1604 and then modified to allow for variability in the input parameters. Other modifications include an increase in the size of the model to accommodate a 16-aircraft squadron, the largest ASW squadron at present, and an explicit form solution to the maintenance queueing equations.

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TABLE OF SYMBOLS

<u>Symbol</u>	<u>Definition or Meaning</u>
a/c	aircraft
λ	mean repair rate of aircraft
λ_A	mean accident rate
S	the set of all possible outcomes; the probability description space
E	possible outcome(s) or event(s)
$p_{ij}(n, n + 1)$	a conditional probability that at time $n + 1$ the outcome or state is j given that at time n the state is i
$X_1(t)$	the number of a/c flying at time t , which did not fly the previous cycle
$X_2(t)$	the number of a/c in the maintenance queue at time t
A(t)	the number of a/c desired on station at time t
N(t)	the total number of a/c of type considered at time t
T	the time interval from the launch to recovery at the start of the cycle
$Q(t_0)$	the probability distribution over all possible states at initial time t_0
P(t)	the matrix of transition probabilities at time t

<u>Symbol</u>	<u>Definition or Meaning</u>
$P(\alpha, i) (\beta, j)$	the elements of the P matrix; the probability that $X_1 = \beta$ and $X_2 = j$ at the end of a cycle, given that $X_1 = \alpha$ and $X_2 = i$ at the start of the cycle
γ_{fgh}	the probability given f ready a/c, g are launched, and h enter maintenance
P_γ	probability of entering maintenance just before, during, or immediately after launch
p	the probability of equipment failure during flight requiring maintenance when recovered by the carrier
$\Pi_\alpha(m)$	the probability that of α a/c flying m will enter maintenance upon recovery
D	the number of independent identical maintenance repair stations or "spots"
$p_{ij}(t)$	the probability that i - j a/c are repaired in time interval t

1. INTRODUCTION

The threat to freedom of the seas posed by the vast Soviet submarine fleet is perhaps the most thorny problem facing the U. S. Navy today. Two world wars have produced Pyrrhic victories over limited submarine fleets. During the Second World War operations analysis was born into the Navy to aid in the defeat of the German submarine. The classic antisubmarine warfare (ASW) analyses and models developed by Morse [2] and Koopmans [3] are still being used today, over two decades later, in most of the ASW study efforts for the Navy.

These early ASW analyses assumed a given level of search effort available and directly evaluated the probability that an ASW subsystem could detect and/or kill a submarine. This assumption is not only logical to make the problem tractable, but also practical since no immediate changes in ASW force levels could be expected. Moreover, the studies were conducted during the war, not before it started. It is the purpose of this paper to present a probabilistic model to describe the available effort. Such a model can be used to sharpen the estimates of the effectiveness of an ASW subsystem and to study the characteristics of the associated support system.

Naturally, the current study plays an important but limited role in the overall problem of designing an entire ASW system. The difficulties involved in such a specification are legion. First and foremost

is the quantification of the ASW mission in denying the enemy the effective use of his submarines. Currently, the probability of detecting and/or killing submarines is used as the measure of effectiveness of the mission, and it appears that a more encompassing one has not been developed. Second, the specification of an ASW force level to counter a given threat has many inherent subjective elements. These are due to the existing historical bias in predicting the conduct of a future ASW war with an enemy, particularly one who has never before used a large submarine force in its military operations. The reader can imagine why merely defining terms such as "threat" and "effective counter" becomes quite difficult.

Thus, there is a need to investigate the levels of search effort specified. This may require acceptable models to measure the availability of effort, its effectiveness, and determine the logistic support required for any level of available effort. Specifically, the ASW subsystem to be modeled is the carrier-based aircraft, although the model is adaptable to other systems.

The method of investigating the demand for ASW carrier a/c will assume that the desired number of a/c on station is known as an input parameter. The support required to achieve this measure of available effort depends upon maintenance space, manpower, and supply. Generally, we shall consider how an ASW carrier supports this number of a/c on station with the present or proposed number of a/c embarked

on the carrier. The parametric input can be subjected to sensitivity analyses.

The operational commander of the ASW force launches the desired number of a/c on station to screen, search, or actively prosecute a submarine contact. Each a/c is relieved on station. Each such relief requires the launching of another a/c prior to the recovery of the initial a/c. The returning a/c must receive varying degrees of maintenance and requires refueling and rearming. This cycle continues until the mission is completed. Loss of a/c due to accidents, insufficient supply, and lack of repair capability cause deviations in this procedure. Naval operations involve the interaction of many quantities which are random in nature. Not all can be considered in a tractable mathematical model. Some quantities which are important are omitted. One example is the length of each cycle time, which is assumed to be a constant value. Including variables of this nature incurs unnecessary mathematical complication. It is hoped that adequacy of the model can be measured by using fleet data available from the Fleet ASW Data Analysis Program (FADAP).

Collins [5] describes a bivariate Markov model for airborne early warning (AEW) and combat air patrol (CAP) jet a/c operating in an attack carrier force. This model is used to evaluate the probability of maintaining a fixed requirement of a/c on station as a measure of effectiveness of the system. It has subsequently been used in a larger

attack force study for the Navy. The model computes the probabilities of the number of a/c on station and in or awaiting maintenance at any given launch period. The comparable ASW problem differs in the following aspects:

1. Type, range, and speed of a/c;
2. The variable number of a/c required for mission;
3. Attrition due to accidents and supply failures;
4. The greater number of ASW a/c.

It was decided to use the Collins' model with appropriate modification. For immediate reference, the mathematical content of the model will be repeated herein.

In order to incorporate these modifications, it was necessary to spend some time reprogramming on the CDC 1604 digital computer in FORTRAN 63, the CDC version of the IBM FORTRAN IV. The original program was not readily available and was written in an early assembler language. Moreover, the numerical analysis was not sufficiently sharp to handle the larger input values. Also, double precision (two computer words instead of one) arithmetic was required in one subroutine for an accurate explicit solution to the maintenance queueing equations (see Appendix I). This effected a 50% decrease in the computer time required for developing a matrix of transition probabilities.

Following this introduction, section 2 contains a brief description of the operational problems involved and the assumptions made. A brief

description of Markov chains and the mathematical model are presented in section 3. The details for computing the matrix of transition probabilities are given in section 4. General employment of the model follows. The appendices include the solution mentioned on the preceding page, a logical flow diagram of the program, a copy of the program, and some sample results.

2. ASSUMPTIONS

The real-world employment of carrier a/c is cyclic in nature, and the present state of any given a/c (i. e. , flying, in or awaiting maintenance) depends largely on what the previous state was. This fact suggests that a Markovian assumption can logically be made for the a/c transition probabilities. In the search phase, a/c may or may not relieve on station; but, in any part of the contact investigation phase, relief on station will be made. To insure full screening and mission coverage, a/c will relieve on station.

The question of resupply during an operation depends primarily on the availability of carrier on-board delivery (COD). This depends on the geographical location and the mission (convoy protection, strike-force protection, hunter-killer operation, etc.). In practice, resupply is not anticipated within a week's period, and around-the-clock operations have continued for two weeks without resupply.

Standard maintenance procedures aboard carriers preclude major maintenance on the flight deck. It will be assumed that sufficient notice is given so that all major 120-hour checks will be completed prior to the operation. This assumption can be modified with an appropriate adjustment in the mean repair rate. The concept of maintenance crews assigned to hangar deck areas ("spots"), as developed by Collins [3], will be used. Each crew will be capable of all types of maintenance

and will operate independently at the identical mean repair rate λ .

The number of spots is determined by the average number of such crews available to work continuously around the clock on a watch basis.

The state of each a/c is assumed to be statistically independent of that of others, and the launching and landing transition probabilities will be developed on the basis of independent Bernoulli trials. The parameters can be determined using the maximum likelihood estimators. The range of the number of a/c desired on station at any given cycle will be set by the user. The number to be launched at any time is assumed equally likely within this range. This input parameter is a function of the estimated submarine density (i. e., expected contact rate). The lower limit will be set at the number of a/c desired on station in the search (screening) phase, and the upper limit is set at the maximum practicable number of a/c to be launched during a multiple-contact phase.

Briefly, the assumptions are:

1. a/c will be relieved on station.
2. Any desired length of operation can be set as an input.
3. Major 120-hour checks will be completed prior to the operation.
4. No resupply to the carrier is available.
5. The launch-to-launch cycle for all ASW a/c is four hours.
6. Minor maintenance, refueling, and rearming only can be performed on the flight deck.

7. Each maintenance spot is characterized with an independent exponential repair time with mean repair rate of λ for around-the-clock operations.
8. The number of a/c lost due to attrition is a Poisson random variable for each cycle period with parameter λ_A (a/c accident/flying hours for a/c type).
9. Any a/c lost by accident will not be returned to service due to either (a) physical loss at sea, or (b) insufficient maintenance capability aboard ship and lack of major parts.
10. The number of a/c launched for each cycle is uniformly distributed between the upper and lower limits determined by the user.

3. MODEL DESCRIPTION

3.1 The Theory

A stochastic or random process is a collection of random variables indexed on some set T , $(X(t), t \in T)$. In this case, time is the indexing set, and the Markovian assumption states that the future state of the process depends only on the state at the present time and not on its past history. Due to the cyclic nature of our problem, it is possible to increment time ($T = (0, 1, \dots)$) using the cycle time from launch to launch as the steps of unit time in a discrete Markov chain. It is assumed that the reader is familiar with the notion of a random variable as a function defined on a sample description space (S) on which the family of events or outcomes (E) of a probability function can be defined [4].

A discrete time Markov chain is described by a sequence of discrete valued random variables and is determined when the one-step transition probabilities of the state variables are specified, i. e., a conditional transition probability of a transition at time n for each pair of $i, j = 0, 1, \dots, m$ (m being the number of states in the process) must be given.

$$p_{ij}(n, n+1) = P [X(n+1) = j \mid X(n) = i]$$

If the transition probability functions depend only on the time difference, we have time homogeneity

$$p_{ij}(n+1, 1) = p_{ij}(0, 1) = p_{ij} .$$

The initial state of the system must be given either as a specific state or randomly as a probability distribution function over the possible states.

The p_{ij} (transition probabilities) are arranged in matrix form and satisfy:

1. $p_{ij} \geq 0$ for $i, j = 0, 1, \dots, m$;
2. $\sum_{j=0}^m p_{ij} = 1$, i. e., the rows of the transition matrix sum to 1 for all i for the states within the description space [4].

3.2 The Model

In order to establish the finite set of states (E) for the model, we shall consider two random variables defined as follows:

$X_1(t)$ = The number of a/c flying at time t not having flown in the previous launch-to-launch interval.

$X_2(t)$ = The number of a/c in or awaiting maintenance at time t .

Now, we will consider the vector $X(t) = [X_1(t), X_2(t)]$ as a pair of random variables and thereby have a bivariate stochastic process with the possible states ranging from $(0, 0)$ to (A, N) .

$0 \leq X_1(t) \leq A$ = No. of a/c desired on station, and

$0 \leq X_2(t) \leq N$ = No. of a/c of given type aboard carrier.

We will define an operating cycle as an interval unit of time. Process observations of $X(t)$ will be made at successive unit interval launch times. To develop the p_{ij} elements, consider a given time t for launching until A aircraft are flying or until the supply of ready a/c is depleted. Those a/c failing the launch enter the maintenance state at this idealized point in time t (the total launching time required is much less than the total cycle time). At some time T , less than the launch-to-launch unit time interval, the a/c which were relieved on station return and land at the idealized point in time $t + T$. Some of these a/c will require maintenance and enter the maintenance queue. Those requiring only refueling and preflight inspection will enter a ready status to be tested for the next launch.

During the unit time interval, maintenance will be performed on those a/c in the not-ready status, and a certain number of aircraft will be repaired according to assumption 7.

In summary, we start the system in some initial state (such as $(0, 0)$ with no a/c flying or in maintenance) or start with a probability distribution $Q(t_0)$ over the states, E , at time t_0 . We launch, recover, and repair a/c in the unit interval and repeat the process over each succeeding unit time interval until the end of the operating period. Knowing the transition probabilities within the unit time interval, we can develop the elements of the transition matrix, P , or $\{p_{(\alpha, i), (\beta, j)}\}$. These are the probabilities of going from the state of α a/c flying and i a/c in maintenance to β a/c flying and j a/c in maintenance over the unit time interval.

It was assumed in section 2 that A , the number of a/c to be launched, and N , the total number of a/c on board, are random variables, whereas they have been treated as constants so far in the development. To be analytically correct in including this feature, one should develop the appropriate quadrivariate process. Such a development leads to too large a state space and the author chose to include these effects by using a Monte Carlo simulation technique. That is, at the beginning of each cycle, a random mechanism is used to determine the values on A and N .

The probability of losing an a/c or changing the desired number to be launched is determined from the specified distributions at the beginning

of each unit interval, and the resulting P matrix containing the $P(\alpha, i), (\beta, j)$ is then recomputed. The probability distribution Q(t) over the states at any time t may be determined by the appropriate number of successive iterations of the Q vector times the P matrix, i. e.,

$$Q(t) = P[X_1(t) = \beta, X_2(t) = j] = Q(t - 1) \times P .$$

The probability of maintaining α a/c on station over any given period of operation may be obtained at any unit time t (i. e., the beginning of the next cycle) by summing out the appropriate maintenance state probabilities. Thus, $P(\alpha$ a/c are flying at time t) =

$$\Pr (X_1(t) = \alpha) = \sum_{i=0}^N \Pr (X_1(t) = \alpha, X_2(t) = i) .$$

A mathematical comment appears to be in order. In the case of fixed A and N, the states of the Markov chain are positive recurrent; and steady-state probabilities can be found for the entire state space. In the case of decreasing N due to a/c attrition, this is not true; and (0, 0) becomes an absorbing state as time (t) goes to infinity. This latter consideration is not a realistic one for the operational period envisioned. Therefore, it is mathematically more feasible to use the former chain in conjunction with the Monte Carlo technique.

4. DEVELOPMENT OF THE TRANSITION MATRIX

Perhaps the simplest way to view this development is to note the various transition probabilities incorporated in one-unit time cycle defined as follows:

- (1) γ_{fgh} = the launching transition probabilities at time t . This is the probability of taking f ready a/c, launching g successfully, and sending h into maintenance. Each a/c to be launched is considered a Bernoulli trial with probability of failure of p_γ , which is estimable and subject to sensitivity analysis. The values of γ_{fgh} are:
- 0 if $g > A$, since only A a/c are desired;
 - 0 if $g + h > f$; it is impossible to launch and send into maintenance more a/c than are available;
 - 0 if $g < A$, $g + h < f$; launching continues until A a/c are flying or until all f are used up;
 - $\binom{f}{g} (1 - p_\gamma)^g (p_\gamma)^{f-g}$ if $g < A$, $g + h = f$, standard binomial when all a/c in the ready state are used up but the A a/c are not launched;
 - $\binom{g+h-1}{h} (1 - p)^g (p)^h$ if $g = A$, $g + h > f$, standard negative binomial for g successes in $g + h - 1$ trials.
- (2) $\Pi_\alpha(m)$ = the landing transition probabilities which occur at time $t + T$. We must consider the probability that if there are a/c flying at time t then m a/c will enter maintenance at recovery time $t + T$.

$\Pi_{\alpha}(m)$ will equal a standard binomial where p = the probability of equipment failure in flight:

$$\Pi_{\alpha}(m) = \binom{\alpha}{m} (1 - p)^{\alpha - m} (p)^m, \quad m = 0, 1, \dots, \alpha.$$

(3) $p_{ij}(\tau)$ = the maintenance transition probabilities, i. e., the probability of repairing $(i - j)$ a/c in time τ . Two maintenance periods occur: the first starting at time t and ending at time $t + T$, the second starting at time $t + T$ and ending at the end of the cycle, $(t + 1)$. Under assumption 7, the pulsed input, multiple exponential server queue is developed with D maintenance "spots" or servers each with identical, independent service rates, λ . For each server, then, the probability of remaining occupied (given the server is busy) in time $\tau = e^{-\lambda\tau}$. The probability of becoming free (i. e., repairing an a/c) = $1 - e^{-\lambda\tau}$. The resulting queueing equations are:

$$A. \quad dP_{i, n}(t) / dt = -n\lambda P_{i, n}(t) + (n+1)\lambda P_{i, n+1}(t) \quad \text{for } 0 \leq n < D;$$

$$B. \quad dP_{i, n}(t) / dt = -D\lambda P_{i, n}(t) + D\lambda P_{i, n+1}(t) \quad \text{for } n \geq D.$$

Three ranges of i (initial queue state), j (final queue state), and D become significant:

a. When $j \leq i \leq D$, then not all spots are busy since there are fewer a/c in maintenance than spots. Each spot works independently; therefore, the solution to A is the binomial:

$$p_{ij}(t) = \binom{i}{j} (1 - e^{-\lambda t})^{(i-j)} e^{-\lambda t j}.$$

- b. When $D \leq j \leq i$, then all spots are occupied throughout the total service time, and the closed form solution to B is the Poisson:

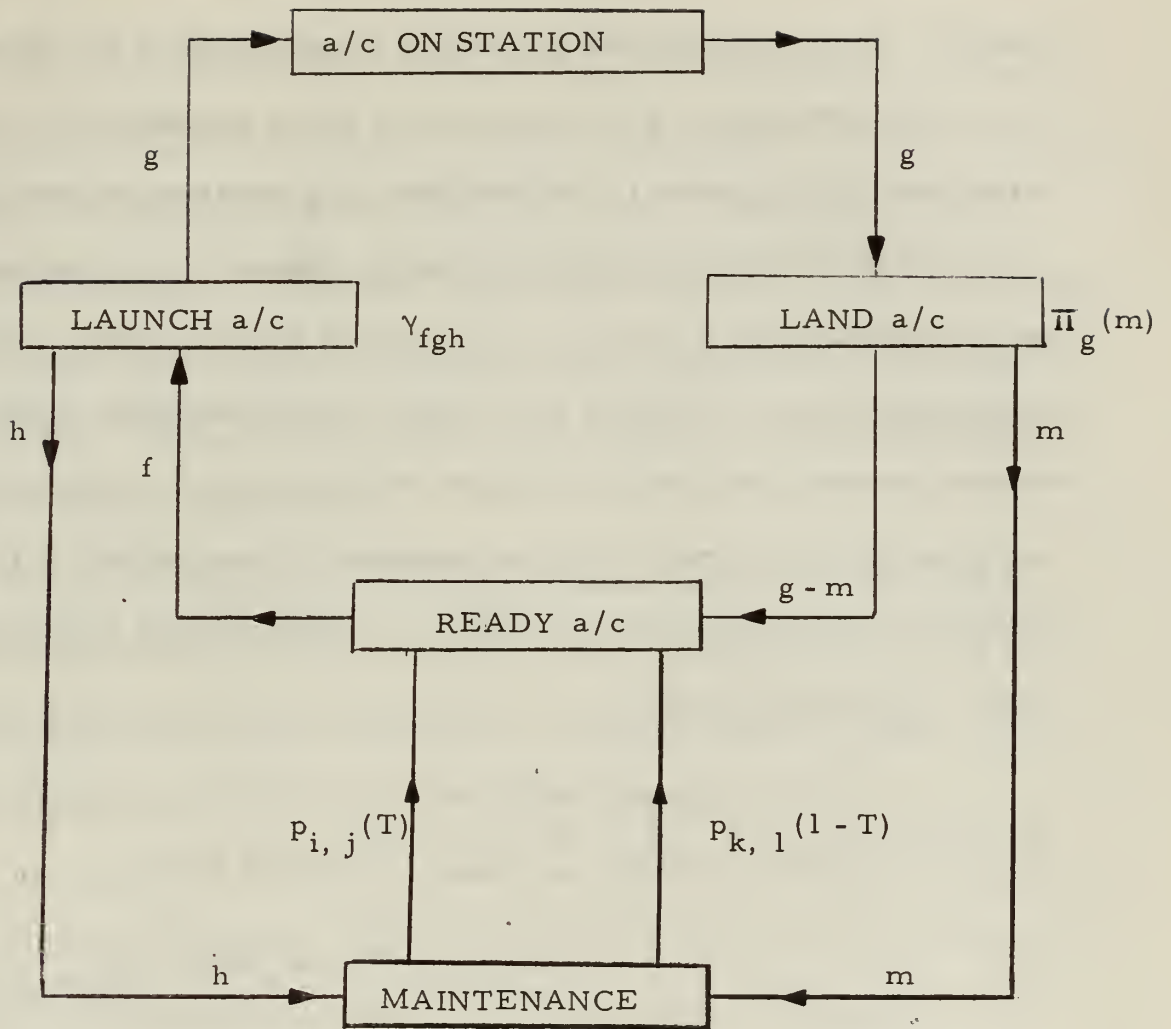
$$P_{ij}(t) = \frac{(D\lambda t)^{(i-j)} e^{-D\lambda t}}{(i-j)!} .$$

- c. When $j < D < i$, then all spots are busy at the beginning of the service period, and some spots become idle during the service period. The explicit form solution of equation A is found using moment generating function transformation:

$$P_{ij}(t) = \sum_{n=j}^{D-1} \binom{n}{j} \binom{D}{n} \left\{ \left(\frac{D}{D-n} \right)^{(i-D)} e^{-\lambda t n} \right. \\ \left. - e^{-\lambda t} \sum_{k=0} \left(\frac{\lambda D t}{k!} \right)^k \left(\frac{D}{D-n} \right)^{i-D-k} \right\} .$$

(The derivation of this solution is discussed in Appendix I.)

The figure on the following page will show the relationships of these transition probabilities within the unit time interval.



TRANSITION PROBABILITIES WITHIN THE UNIT CYCLE

FIGURE 1

In order to develop each transition probability over the total unit time interval, we must consider all events taking place within the interval. Thus, to obtain the probability of going from α a/c flying and i a/c in maintenance to β a/c flying and j a/c in maintenance, we start at the state (α, i) at time t . At this time, α a/c are launched and some l a/c failing the launch enter maintenance. These $i + l$ in maintenance are then serviced until time $t + T$ when some k a/c are still in the maintenance state. At time $t + T$, of the α a/c previously flying, some m enter maintenance and $(\alpha - m)$ enter the ready pool. Maintenance is continued on the $(k + m)$ a/c for the remainder of the cycle $(1 - T)$, until the end of the unit time interval when j a/c remain in the maintenance state. In functional form:

$$P(\alpha, i), (\beta, j) = \sum_{l=0}^{N-\alpha-i} \sum_{k=0}^{i+l} \sum_{m=0}^{\alpha} \gamma_{N-\alpha-i, \beta, l} \cdot P_{i+l, k}(T) \cdot \Pi_{\alpha}(m) \cdot p_{k+m, j}(1-T)$$

5. SUMMARY

Representative values for the mean repair rate and the landing and launching failure rates produced results in agreement with the sensitivity analysis by Collins on these parameters in [5]. For failure probabilities less than .5, and mean repair rate less than 12 hours, the effect of reducing the available maintenance time to 80% of the cycle time was negligible. Optimal loading and cycling policies can be determined for known values of these rates.

The model affords the following checks: (1) the rows of each P matrix are summed as they are computed by the program; and (2) the probability distribution vector (QJ) is summed over the states. Each summation was within 10^{-8} of one in the computer model.

The user may substitute any available distribution over the interval of a/c desired on station. In order to keep A fixed, enter the desired value as both upper and lower limit (A = ALOLIM = LUPLIM). For fixed N, use a very small value for ALAM (such as 10^{-8}). Subroutine KRAN is a uniform generator, using the half open interval (lower limit + 1, upper limit + 2) and a starting number as inputs. KRAN outputs an integer in this interval. Subroutine DRAW was used to provide some intuitive grasp of the results. DRAW was used in binary card form and is not essential to the main program. (The indicated associated statements must be removed, however.)

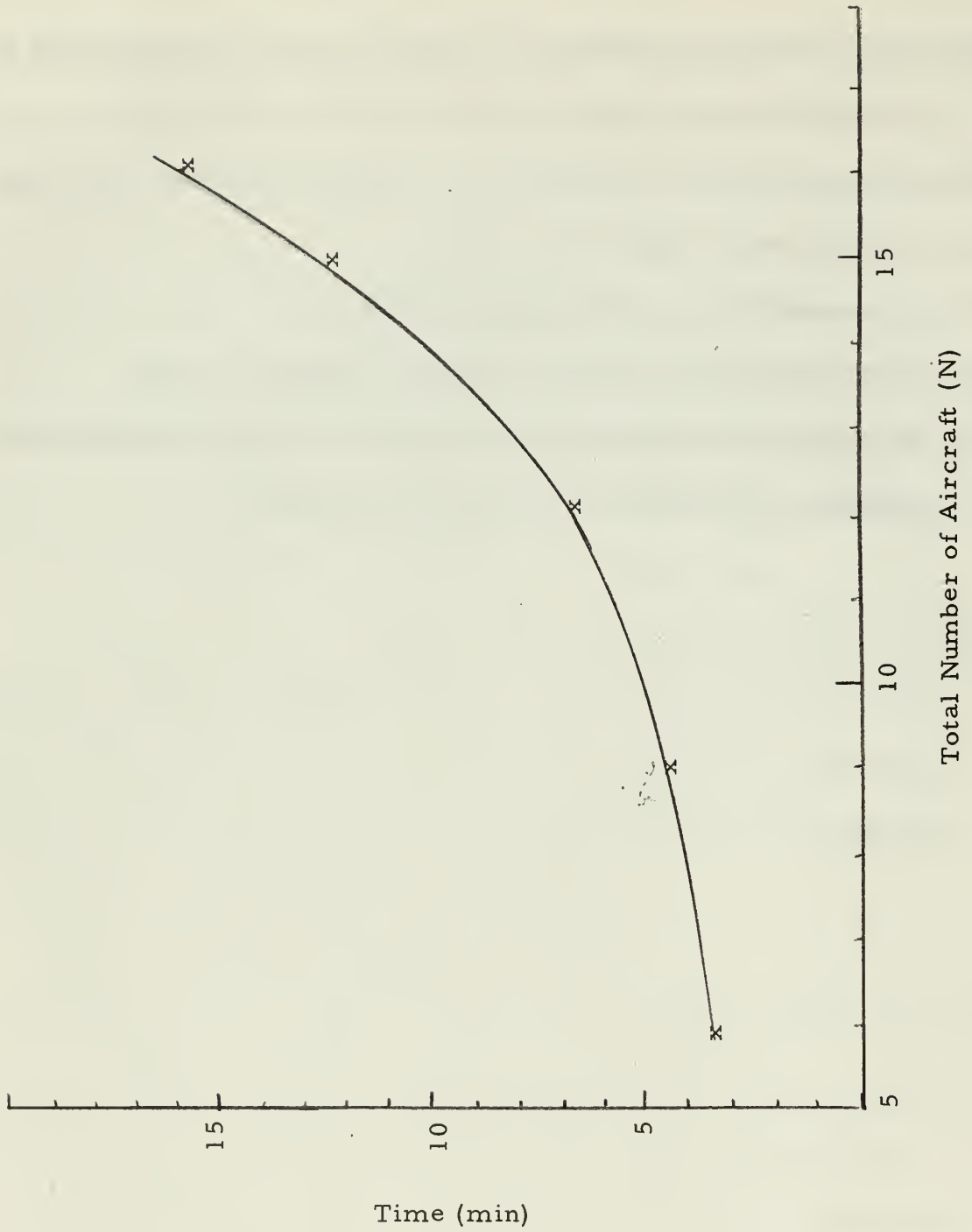
The results of reasonable arbitrary parameter values, based on the author's experience, have shown that most of the probabilities concentrate over a few states. Moreover, computation time increases rapidly as a function of N (no. of a/c), see Figure 2. This would indicate that a simple approximation to the model could be developed. One method presently being investigated to reduce computation time is to shrink the probability state space to include only those significant states and, thus, reduce the size of the transition matrix. Alternatively, the eigenvector, eigenvalue representation of the P matrix, might be used.

Originally, it was hoped to utilize the data from the Fleet ASW Data Analysis Program (FADAP) to attempt a verification of the model with its real-world counterpart. The only method available at present for obtaining the necessary data is by direct observation or a program of data collection, as suggested by Collins [5] .

Many fruitful areas of investigation exist:

- (1) Attrition has been simply modeled by the Poisson method. The two components of attrition, accidents and supply shortage, can be more accurately modeled and used to develop logistic schedules for maintenance and supply. One simple technique is to assume each component is independent and Poisson, and estimate a supply failure rate for AOCF attrition from past data. With these assumptions, the total attrition is Poisson, with the parameter equal to the sum of the accident and supply failure rates.

- (2) The model can be modified to make the number of maintenance spots available for any cycle a variable function of time, $D(t)$.
- (3) An investigation of the Markovian assumption validity as the cycle times become smaller and smaller.
- (4) Development of a continuous time model.
- (5) Modification of the model to simulate resupply by COD.
- (6) A study of the distribution of submarine contacts to determine the validity of the uniform a/c demand assumption.



PROGRAM ASSEMBLY AND COMPUTATION TIME
 FOR ONE TRANSITION MATRIX (P) AS A FUNCTION
 OF THE TOTAL NUMBER OF AIRCRAFT (N)

FIGURE 2

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APPENDIX I

EXPLICIT SOLUTIONS OF THE MAINTENANCE QUEUEING EQUATIONS

The queueing equations for the pulsed input queue are essentially the pure death process given in [1] and [4] as problems and developed by Collins in [5]. The equations are:

$$A. \quad \frac{dP_{i, n}(t)}{dt} = -n\lambda P_{i, n}(t) + (n+1)\lambda P_{i, n+1}(t) \quad \text{for } 0 < n < D$$

$$B. \quad \frac{dP_{i, n}(t)}{dt} = -n\lambda P_{i, n}(t) + D\lambda P_{i, n+1}(t) \quad \text{for } n \geq D$$

where $P_{ij}(0) = \Delta_{ij}$ and $P_{ij}(t) = 0$ for $i < j$, since no input (arrivals) occur during the service time.

Equation B is solved directly in closed form:

$$P_{i, n}(t) = \frac{(D\lambda t)^{(n-i)} e^{-\lambda Dt}}{(n-i)!} .$$

Now transforming the first equation (A) using the moment generating function (MGF),

$$G(s, t) = \sum_{n=0}^{D-1} s^n P_n(t) ,$$

as outlined in [4] (Chapter 7), and its partial derivatives:

$$(1) \quad \frac{dG}{dt} = \sum_{n=0}^{D-1} s^n P'_n(t)$$

$$(2) \quad \frac{dG}{ds} = \sum_{n=0}^{D-1} n s^{n-1} P_n(t)$$

Where $P_n(t)$ denotes the conditional probability $P_{i,n}(t)$, by substituting (A) into (1), properly identifying the first summation with (2), and changing the second summation index to $r = n + 1$, we get:

$$\frac{dG}{dt} = -\lambda s \frac{dG}{ds} + \lambda \sum_{r=0}^D r s^{r-1} P_r(t), \quad \text{or}$$

$$(3) \quad \frac{dG}{dt} = -\lambda (s-1) \frac{dG}{ds} + \lambda D s^{D-1} P_D(t),$$

since

$$\sum_{r=0}^D r s^{r-1} P_r(t) = \frac{dG}{ds} + D s^{D-1} P_D(t).$$

Next, replace the partial differential equation (3) with a system of ordinary differential equations using the Lagrangian auxiliary equations:

$$\frac{dt}{1} = \frac{ds}{\lambda (s-1)} = - \frac{dz}{\lambda D s^{D-1} P_D(t)}.$$

The solution to the first equation (using the first two differentials) is:

$$\lambda t = \ln (s-1) + C'$$

and hence

$$s = C_1 e^{\lambda t} + 1$$

or

$$G_1 = e^{-\lambda t} (s - 1) .$$

The second equation is: (using first and third differentials)

$$dz = -\lambda D (C_1 e^{\lambda t} + 1)^{D-1} P_D(t) dt .$$

Using the solution to (B) where $m = i - D$ to replace $P_D(t)$ and integrating, term wise, the binomial expansion of $(C_1 e^{\lambda t} + 1)^{D-1}$:

$$z = \frac{(\lambda D)^{m+1}}{m!} \sum_{j=0}^{D-1} \binom{D-1}{j} C_1^j \int t^m e^{-\lambda(D-j)t} dt$$

where the integral is evaluated as:

$$- \sum_{k=0}^m \frac{t^k e^{-\lambda(D-j)t}}{(\lambda(D-j))^{m-k+1}} \frac{m!}{k!} + C_2 .$$

Thus,

$$C_2 = z + e^{-\lambda Dt} \sum_{j=0}^{D-1} \binom{D}{j} (s-1)^j \sum_{k=0}^m \frac{(\lambda Dt)^k}{k!} \left(\frac{D}{D-j} \right)^{m-k}$$

and the general solution is $\phi(C_1, C_2)$, where ϕ is an arbitrary function

and

$$C_1 = u(s, t, z)$$

and

$$C_2 = v(s, t, z) .$$

To get our particular solution, use the boundary conditions for $G(s, t)$:

(1) for $s = 1$,

$$G(1, t) = \sum_{n=0}^{D-1} P_n(t)$$

$$= \Pr [\text{no. in maintenance at } t \text{ is } < D \mid i \text{ at } t = 0]$$

$$G(1, t) = 1 - \sum_{n=0}^{i-D} \frac{e^{-\lambda Dt} (\lambda Dt)^n}{n!} = 1 - \psi_1(t)$$

where

$$u(1, t, z) = C_1 = 0$$

$$v(1, t, z) = C_2 = z + e^{-\lambda Dt} \sum_{k=0}^m \frac{(\lambda Dt)^k}{k!}$$

so

$$C_2 = z + \psi_1(t)$$

(2) for $t = 0$,

$$G(s, 0) = \sum_{n=0}^{D-1} s^n P_n(0) = 0, \quad \text{since } i \geq n > D$$

where

$$u(s, 0, z) = (s - 1)$$

$$v(s, 0, z) = C_2 = z + \sum_{j=0}^{D-1} \binom{D}{j} (s - 1)^j \left(\frac{D}{D - j} \right)^m .$$

Thus,

$$G(s, 0) = z + \sum_{j=0}^{D-1} \binom{D}{j} C^j \left(\frac{D}{D - j} \right)^m - C_2 .$$

Substituting the general value for C_2 above:

$$G(s, t) = \phi(u, v) = \sum_{j=0}^{D-1} \binom{D}{j} (s-1)^j e^{-\lambda t j} \left(\frac{D}{D-j}\right)^m$$

$$- \sum_{j=0}^{D-1} \binom{D}{j} (s-1)^j \sum_{k=0}^m \frac{(\lambda Dt)^k}{k!} e^{-\lambda Dt} \left(\frac{D}{D-j}\right)^{m-k}$$

Rearranging terms,

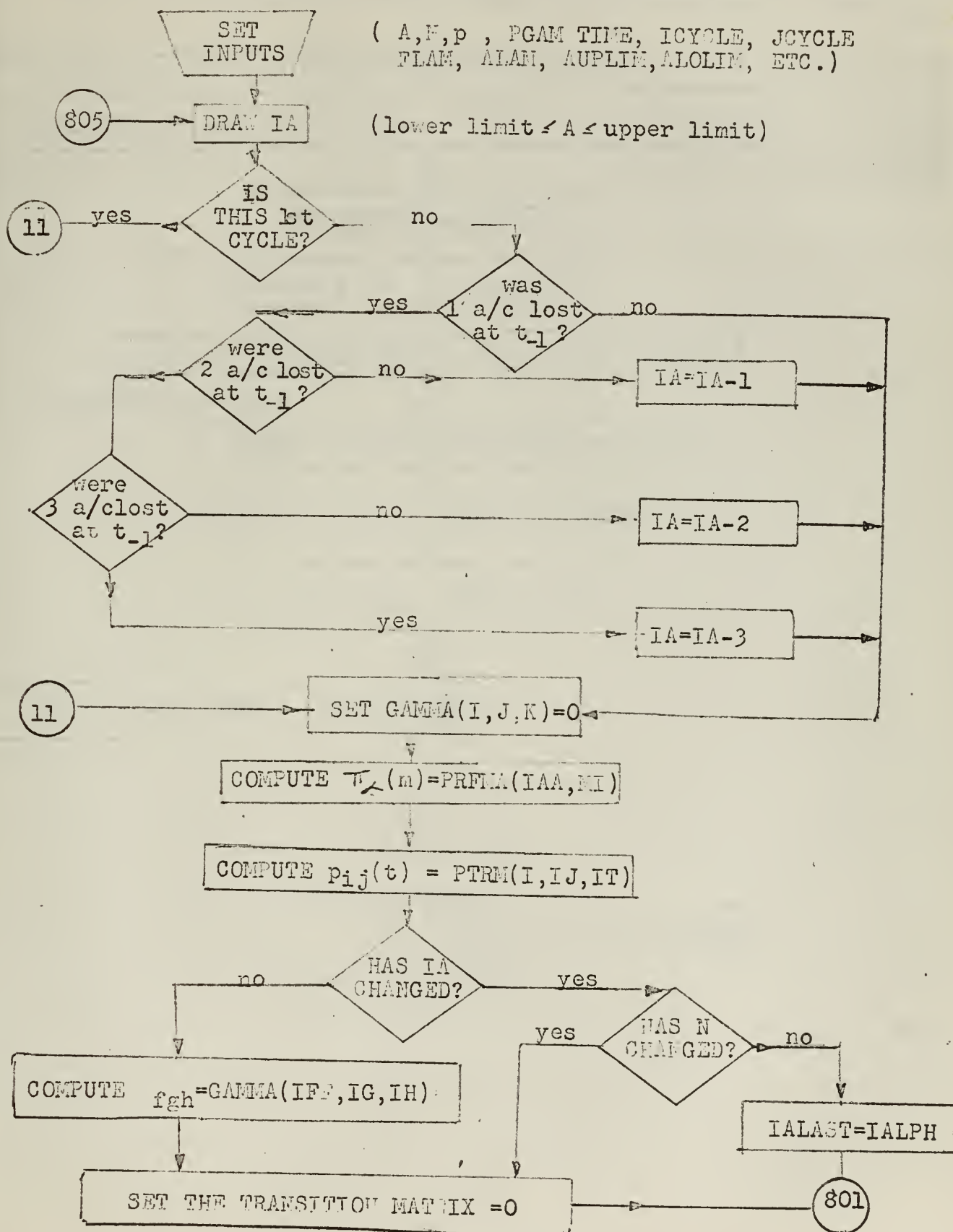
$$G(s, t) = \sum_{n=0}^{D-1} s^n \sum_{j=n}^{D-1} \binom{j}{n} \binom{D}{j} (-1)^j \left[\left(\frac{D}{D-j}\right)^m e^{-\lambda t j} \right.$$

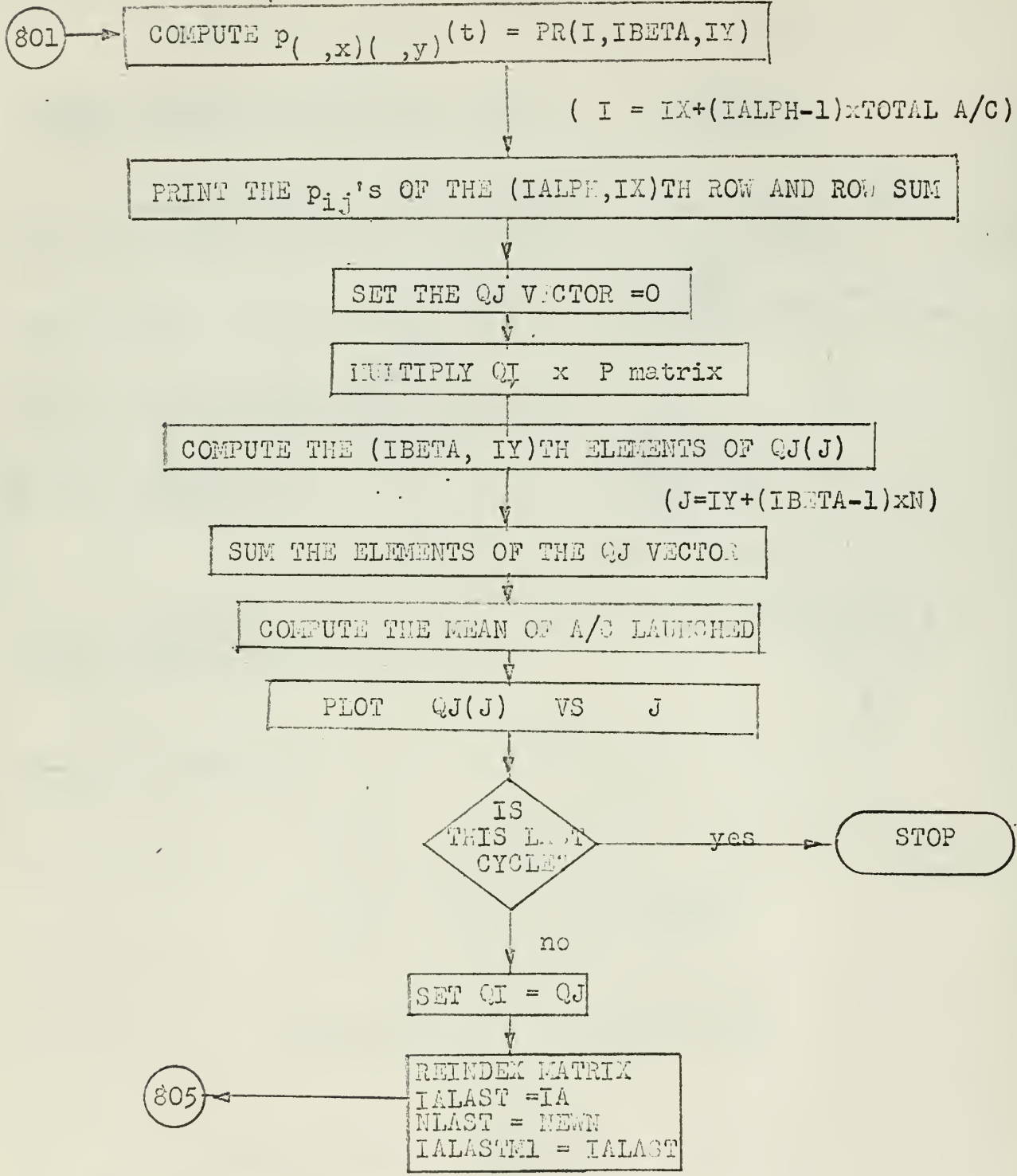
$$\left. - e^{-\lambda Dt} \sum_{k=0}^m \frac{(\lambda Dt)^k}{k!} \left(\frac{D}{D-j}\right)^{m-k} \right]$$

where $P_n(t) =$ the coefficient of s^n .

APPENDIX II

THE LOGICAL FLOW DIAGRAM OF THE COMPUTER PROGRAM





APPENDIX III

THE COMPUTER PROGRAM

```

-COOP,,LANMAN,0/49/S/1S/2S/E/45=54,15 ,30000,5.          000
-BINARY,56.                                               000
(RELOCOM.                                               000
-FTN,E.                                                 000
      PROGRAM MARKOV                                     000
C THIS PROGRAM IS A NONSTATIONARY BIVARIATE MARKOV CHAIN MODEL OF ASW A/C 000
C OPERATIONS. THE RANDOM VARIABLES ARE THE NUMBER OF A/C FLYING AT THE 000
C BEGINNING OF ANY GIVEN LAUNCH CYCLE. THE MAXIMUM NO. OF A/C ALLOWED IN 000
C THE MODEL IS 16(NA). THE RANGE OF A/C TO BE LAUNCHED AT ANY GIVEN 000
C INTERVAL IS 0 TO 6 A/C. THE FOLLOWING INPUTS ARE REQUIRED. 000
      ID= THE NO. OF INDEPENDENT MAINTENANCE SPOTS       001
      NA= TOTAL NO. OF A/C TYPE ON BOARD                 001
      TIME=TIME FROM LAUNCH TO RECOVERY/LAUNCH TO LAUNCH CYCLE TIME(HRS) 001
      FLAM=MEAN REPAIR TIME PER SPOT/LAUNCH TO LAUNCH CYCLE TIME (HRS) 001
      PGAM= PROBABILITY OF A/C FAILING LAUNCH(M.L. EST. FROM PAST DATA) 001
      P= PROBABILITY OF A/C FAILURE DURING FLIGHT REQUIRING MAINTENANCE 001
          AT LANDING (M.L. ESTIMATOR FROM PAST DATA)    001
      QI= THE PROBABILITY DISTRIBUTION VECTOR OVER ALL POSSIBLE STATES 001
          (7 X 17 = 119) SUCH THAT THE SUM OF ALL QI(I) = 1. THIS 001
          IS ESTIMATED BY THE USER AND INPUTTED BY USING A DATA STATEMENT 001
      ICYCLE = NO. CYCLES DESIRED FOR OPERATION           002
      JCYCLE = LAUNCH TO LAUNCH TIME(HRS)(TOT. TIME=ICYCLE X JCYCLE) 002
      ALAM = ACCIDENT RATE FOR TYPE A/C (ACCIDENT/HOURS)' 002
      ALOLIM = DESIRED LOWER LIMIT ON A                  002
      AUPLIM = DESIRED UPPER LIMIT ON A                  002
      COMMON FLAM,TIME                                   002
      TYPE DOUBLE FLAM                                   002
      COMMON PTRM,GAMMA,PR,PRFMA,ID                     002
      DIMENSION BC(17),A(17),FBC(17)                   002
      DIMENSION PTRM(17,17,2),GAMMA(17,7,17),PRFMA(7,7) 002
      DIMENSION PR(119,7,17),QI(119),QJ(119)           002
      DIMENSION FJPLOT(119),JT(12)                     002
C ENTER DATA CARDS HERE                                002
      DATA((QI(I),I=1,119)=.2,16(.05),102(.0))        003
      NA=16
      ALAM=.01
      ID = 8
      FLAM=3.0
      PGAM=P=.4
      IYY = 13421773
      TIME = .125
      ICYCLE=20
      JCYCLE=4
      ALOLIM=4.
      AUPLIM=6.
      END OF DATA CARDS                                003
      AL=ALOLIM+1. $ AU=AUPLIM +2.                      003
      UNITT=1.                                          003
      N=NA+1                                           003
      IAMAX=7                                           003
      IALAST=0                                           003
      D=FLOATF(ID)                                       003
      NLAST=NEWN=N                                       003
      KT=1                                              003
809 IA=KTRAN(AL,AU,IYY)                                 003
      IF(KT.EQ.1) 113,115                               004

```



```

115 T1=-LOGF(.000000001 + RANF(-1))*2.30258/ALAM 0042
    IF(T1-TFLC)130,131,132 0043
130 T2=-LOGF(.000000001 + RANF(-1))*2.30258/ALAM 0044
    IF(T1+T2-TFLC) 230,231,131 0045
230 T3=-LOGF(.000000001 + RANF(-1))*2.30258/ALAM 0046
    IF(T1+T2+T3-TFLC) 331,331,231 0047
331 NEWN=NLAST-3 $ GO TO113 0048
231 NEWN=NLAST-2 $ GO TO113 0049
132 NEWN=NLAST $ GO TO113 0050
131 NEWN=NLAST-1 0051
113 PRINT 8882,IA,NEWN 0052
    IF(NEWN-IA) 15,13,13 0053
    15 IA=NEWN 0054
    13 IF(IALAST) 11,12,11 0055
    12 CONTINUE 0056
FROM THIS NEXT STATEMENT TO NO. 483 IS CONCERNED ONLY WITH THE GRAPH 0057
DO 482 I=1,12 0058
482 JT(I)=8H 0059
    JT(1)=8HE(A/C) = 0060
    JT(3)=8HSPOTS = 0061
    JT(5)=8H T = 0062
    JT(7)=8HJ VS QJ 0063
    JT(8)=8HVECTOR 0064
    JT(9)=8H N = 0065
    JT(11)=8H A = 0066
    DO 483 I=1,119 0067
    FI=I 0068
483 FJPLOT(I)=FI 0069
    IALAST=IA 0070
    DO 1235 I=1,17 0071
    DO 1235 J=1,IAMAX 0072
    DO 1235 K=1,17 0073
1235 GAMMA(I,J,K)=0.0 0074
    AT THIS PT THE LANDING TRANSITION PROBABILITIES ARE COMPUTED 0075
    DO 300 IAA=1,IAMAX 0076
    DO 301 MI=1,IAMAX 0077
    IF(IAA-MI)31,32,33 0078
    31 PRFMA(IAA,MI)=0. 0079
    GO TO 301 0080
    32 MM1=MI-1 0081
    PRFMA(IAA,MI)=P**MM1 0082
    GO TO 301 0083
    33 IAM1=IAA-1 0084
    MM1=MI-1 0085
    BC(1)=1.0 0086
    PROD=FLOATF(IAA-MI) 0087
    DO 50 IP=2,MI 0088
    AIP=FLOATF(IP-1) 0089
    PROD = PROD +1.0 0090
    50 BC(IP)=PROD*BC(IP-1)/AIP 0091
    IGO=IAA-MI 0092
    PRFMA(IAA,MI)= (BC(MI)*(1.0-P)**(IGO))*P**MM1 0093
301 CONTINUE 0094
300 CONTINUE 0095
    AT THIS PT THE MAINTENANCE TRANSITION PROBABILITIES ARE COMPUTED 0096
    DO 100 IT=1,2 0097

```

```

      IF(IT-1)25,25,26
25  TAU = TIME
      GO TO 28
26  TAU = UNITT-TIME
28  DO101 I=1,N
      DO 102 IJ=1,N
          IF (I-IJ) 14,199,17
199  IF(I-ID) 19,19,1999
1999 PTRM(I,IJ,IT)=EXPF(-FLAM*TAU*D)
      GO TO 102
14  PTRM(I,IJ,IT)=0.
      GO TO 102
19  FJM1=FLOATF(IJ-1)
      PTRM(I,IJ,IT)=EXPF(-FLAM*TAU*FJM1)
      GO TO 102
17  IF(I-ID-1) 1,1,2
      1 BC(1)=1.0
          PROD=FLOATF(I-IJ)
          DO 10 IP =2,IJ
              AIP=FLOATF(IP-1)
              PROD = PROD + 1.0
10  BC(IP) =PROD*BC(IP-1)/AIP
          ELT=EXPF(-FLAM*TAU)
          PTRM (I,IJ,IT)=BC(IJ)*(1.-ELT)**(I-IJ)*ELT**(IJ-1)
          GO TO 102
      2 IF(IJ-1-ID) 22,24,24
22  CONTINUE
      CALL PID(I,IJ,IT)
      GO TO 102
24  D=FLOATF(ID)
      ELDT=EXPF(-D*FLAM*TAU)
      FACT = 1.0
      A(1)=1.0
      MM = I-IJ
      DO 20 M=2,MM
          FACT=FACT+1.0
20  A(M)=A(M-1)*FACT
201 PTRM(I,IJ,IT)=(D*FLAM*TAU)**(I-IJ)*ELDT/A(I-IJ)
102 CONTINUE
101 CONTINUE
100 CONTINUE
      GO TO 120
11  CONTINUE
      IF(IA-IALAST) 120,121,120
121 IF(NEWN-NLAST)111,117,111
117 IALPH=IALASTM1 $ GO TO 801
120 CONTINUE
AT THIS POINT THE LAUNCHING TRANSITION PROBABILITIES ARE COMPUTED
DO 204 IFF=1,N
      IGM = XMINOF (IA,IFF)
DO 203 IG=1,IGM
      IGM1=IG-1
DO 202 IH=1,N
      IHM1=IH-1
      BPROD=((1.-PGAM)**IGM1)*(PGAM**IHM1)
86  IF(IG-IA) 91,87,84

```

91	IF(IG+IHM1-IFF) 84,82,84	0154
87	IF(IG+IHM1-IFF)85,85,84	0155
84	GAMMA(IFF,IG,IH)=0.	0156
	GO TO 202	0157
82	BC(1)=1.0	0158
	PROD=FLOATF(IFF-IG)	0159
	DO 30 IP=2,IG	0160
	AIP=FLOATF(IP-1)	0161
	PROD = PROD + 1.0	0162
30	BC(IP)=PROD * BC(IP-1)/AIP	0163
	IHM1=IH-1	0164
	TEMP= PGAM**IHM1	0165
	TEMP1=(1.-PGAM)**IGM1	0166
	BPROD = TEMP*TEMP1	0167
	GAMMA(IFF,IG,IH)=BC(IG)*BPROD	0168
	GO TO 202	0169
85	FBC(1)=1.0	0170
	PROD=FLOATF(IGM1-1)	0171
	DO 40 IP=2,IH	0172
	AIP = FLOATF(IP-1)	0173
	PROD = PROD +1.0	0174
40	FBC(IP)=PROD*FBC(IP-1)/AIP	0175
	GAMMA(IFF,IG,IH)=FBC(IH)*BPROD	0176
202	CONTINUE	0177
203	CONTINUE	0178
204	CONTINUE	0179
	REMOVE CARDS FROM HERE TO NO 999 IF PRINT OUT NOT DESIRED	0180
	PRINT 9,(((I,IJ,IT,PTRM(I,IJ,IT),IT=1,2),IJ=1,N),I=1,N)	0181
9	FORMAT (1H1/(2(6H PTRM(I2,1H,I2,1H,I2,3H) = E14.5)))	0182
	PRINT 99,(((IFF,IG,IH,GAMMA(IFF,IG,IH),IFF=1,N),IG=1,IA),IH=1,N)	0183
99	FORMAT(1H1/(2(7H GAMMA(I2,1H,I2,1H,I2,3H) = E14.5)))	0184
	PRINT 999,(((IAA,MI,PRFMA(IAA,MI),IAA=1,IAMAX),MI=1,IAMAX)	0185
999	FORMAT(1H1/(2(7H PRFMA(I2,1H,I2,3H) = E14.5)))	0186
	NOW THE TRANSITION MATRIX MUST BE ZEROED	0187
111	CONTINUE	0188
	DO 899 J=1,119	0189
	DO 899 K=1,7	0190
	DO 899 L=1,17	0191
899	PR(J,K,L)=0.0	0192
	START COMPUTING THE ELEMENTS OF EACH ROW, I=IX+ (ALPHA - 1) X TOTAL A/C	0193
	DO 1000 IALPH=1, IALAST	0194
801	CONTINUE	0195
	DO 1100 IX=1,NLAST	0196
	COMPUTE THE P ELEMENTS OF THE IAPH,IX ROW AND SUM THE ROW	0197
	TSUM=0.	0198
	I=IX+(IALPH-1)*N	0199
	DO 800 IBETA=1,IA	0200
	RSUM=0.0	0201
	DO 900 IY=1,NEWN	0202
	PR(I,IBETA,IY)=0.	0203
	ILIM=NEWN-IALPH-IX+2	0204
	PSUM=0.0	0205
	SUM=0.0	0206
	SUML=0.0	0207
	DO 500 IL=1,ILIM	0208
	KLIM=IX+IL-1	0209

	IXPIL = IX + IL - 1	021
	SUMM=0.	021
	DO 600 MI=1,IALPH	021
	SUMK=0.	021
	DO 700 IK=1,KLIM	021
	IKPMI = IK + MI - 1	021
	IF(IXPIL-NEWN) 701,701,700	021
701	IF(IKPMI-NEWN) 702,702,700	021
702	GAMH=GAMMA(ILIM,IBETA,IL)	021
	PTRMH1 = PTRM(IXPIL,IK,1)	022
	PRFMAH = PRFMA(IALPH,MI)	022
	PTRMH2 = PTRM(IKPMI,IY,2)	022
	SUM = GAMH * PTRMH1 * PRFMAH * PTRMH2	022
	SUMK=SUMK+SUM	022
	PSUM=PSUM+SUM	022
700	CONTINUE	022
	SUMM = SUMM + SUMK	022
600	CONTINUE	022
	SUML = SUML + SUMM	022
	PSUM2 =SUML	022
500	CONTINUE	023
	RSUM=RSUM+PSUM	023
	PR(I,IBETA,IY)=PSUM	023
900	CONTINUE	023
	TSUM=TSUM+RSUM	023
800	CONTINUE	023
	PRINT 888 , TSUM,IALPH,IX	023
888	FORMAT (7H TSUM = E15.8,2I5)	023
1100	CONTINUE	023
1000	CONTINUE	023
C	REMOVE CARD FROM HERE TO 889 IF P MATRIX PRINT OUT NOT DESIRED	024
	DO 889 J=1,17	024
	DO 889 K=1,7	024
	DO 889 L=1,17	024
	I=J+(K-1)*N	024
889	PRINT 890,(PR(I,LP,L),LP=1,IAMAX),K,J,L	024
890	FORMAT(7E14.5,2HJ=I2,5HK=1,A,2HL=I2)	024
	DO 898 I=1,119	024
898	QJ(I)=0.0	024
C	NOW MULTIPLY QI AND P TO GET QJ	024
805	PRINT 807,KT,IALAST,IA	025
807	FORMAT(1H1,13HQ VECTOR CASE I3/// I5,I5)	025
	DO 802 IBETA=1,7	025
	DO 902 IY=1,17	025
CAT	THIS POINT CALCULATE THE (IBETA,IY)TH ELEMENT OF THE QJ VECTOR	025
	J=IY+(IBETA-1)*N	025
	QP1=0.	025
	QP=0.	025
	DO 2001 IALPH=1,7	025
	DO 2201 IX=1,17	025
	I=IX+(IALPH-1)*N	025
	QP1=QI(I)*PR(I,IBETA,IY)	026
	QP=QP+QP1	026
2201	CONTINUE	026
2001	CONTINUE	026
	QJ(J)=QP	026

PRINT 8882,IBETA,IY,J,QP	0266
882 FORMAT(2I4,4H QJ(I3,3H)= E14.8)	0267
902 CONTINUE	0268
802 CONTINUE	0269
CHECK THE SUM OF THE Q VECTOR	0270
QSUM=0.	0271
DO 808 J=1,119	0272
808 QSUM=QJ(J)+QSUM	0273
PRINT 8883,QSUM	0274
8883 FORMAT(6H QSUM= E15.9)	0275
DO 333 I = 18,119	0276
K = (I-1)/17	0277
FK=FLOATF(K)	0278
FMEAN= FK*QJ(I)+FMEAN	0279
333 CONTINUE	0280
TFLC=FMEAN*FLOATF(JCYCLE)	0281
PRINT 335,FMEAN	0282
335 FORMAT(17HMEAN A/C FLYING = E10.4)	0283
STATEMENTS FROM THIS POINT TO THE CALL DRAW STATEMENT REFER TO GRAPH	0284
JT(2)=ICODE(FMEAN)	0285
JT(4)=ICODE(D)	0286
FKT=FLOATF(KT)	0287
JT(6)=ICODE(FKT)	0288
FN=FLOATF(NEWN-1)	0289
JT(10)=ICODE(FN)	0290
FIAA=FLOATF(IA-1)	0291
JT(12)=ICODE(FIAA)	0292
CALL DRAW(119,FJPLOT,QJ,0,0,4H ,JT,0,0,0,0,0,0,8,8,0,LAST)	0293
FMEAN = 0.	0294
NEXT WE MUST MULTIPLY QJ AND P TO GET QK AND SO ON...(QK+...N)	0295
KT=KT+1	0296
IF(KT-ICYCLE) 803,803,806	0297
803 DO 804 I=1,119	0298
804 QI(I)=QJ(I)	0299
IALASTM1=IALAST	0300
IALAST=IA	0301
NLAST=NEWN	0302
GO TO 809	0303
806 STOP 06	0304
END	0305
SUBROUTINE PID(I,J,IT)	0306
COMMON FLAM,TIME	0307
COMMON PTRM,GAMMA,PR,PRFMA,ID	0308
TYPE DOUBLE BC,BDC,PROD ,DID3,DID4,DID5,DEXP	0309
TYPE DOUBLE DAN,DID1,DID2,SUM,DN,ANM1,FAC,COF,PSUM,PTR,FLAM,TAU,D	0310
DIMENSION PTRM(17,17,2),BC(11),BDC(11)	0311
DIMENSION GAMMA(17,7,17), PRFMA(7,7), PR(119,7,17)	0312
D=FLOATF(ID)	0313
IDP1=ID+1	0314
IF(IT-1)25,25,26	0315
25 TAU = TIME	0316
GO TO 28	0317
26 TAU= 1.-TIME	0318
28 CONTINUE	0319
IMDP1=I-ID	0320
PTR=0.0	0321

```

PSUM=0.
DO 200 NJ = J, ID
C DEVELOP N TAKEN J AT A TIME AND D TAKEN N AT A TIME
BC(1)=1.0
PROD=FLOATF(NJ-J)
DO 10 IP=2,J
AIP=FLOATF(IP-1)
PROD = PROD+1.0
10 BC(IP)=PROD* BC(IP-1)/AIP
BDC(1)=1.0
PROD=FLOATF(IDP1-NJ)
DO 20 IQ=2,NJ
AIQ=FLOATF(IQ-1)
PROD=PROD+1.0
20 BDC(IQ)=PROD*BDC(IQ-1)/AIQ
COF=BC(J)*BDC(NJ)*(-1)**(NJ-J)
ANM1=FLOATF(NJ-1)
DAN=D/(D-ANM1)
DID4=DEXP(-FLAM*TAU*ANM1)
DID1=(DAN**(I-IDP1))*DID4
SUM=0.
DN=0.
DO 201 K=1,IMDP1
FAC=1.
KM1=K-1
PROD=0.
DO 11 IK=1,KM1
PROD=PROD+1.
11 FAC=FAC*PROD
IMIDK=I-ID-K
SUM=((FLAM*D*TAU)**KM1)*DAN**IMIDK / FAC
201 DN=DN+SUM
DID3=DEXP(-FLAM*D*TAU)
DID2=DN*DID3
DID5=DID1-DID2
PSUM=COF*DID5
200 PTR =PTR +PSUM
103 CONTINUE
PTRM(I,J,IT)=PTR
102 CONTINUE
101 CONTINUE
END
FUNCTION KRAN(A,B,IY)
C
C THIS ROUTINE RETURNS AN UNIFORMLY DISTRIBUTED RANDOM INTEGER
C
C THIS ROUTINE RETURNS A INTEGER RANDOM NUMBER .GE. TO A
C .LT. B
C A = BOTTOM LIMIT (INCLUDED) FOR THE RANDOM NUMBER
C B = TOP LIMIT (NOT INCLUDED) FOR THE RANDOM NUMBER
C SET IY ONLY ONCE IN MAIN PROGRAM FOR EACH SET OF RANDOM NUMBERS
C SOME GOOD STARTING VALUES FOR IY FOLLOW
C 13421773
C 33554433
C 8426219
C 42758321

```

56237485

62104023

ANY OF THESE MAY BE USED

THIS ROUTINE MAY BE USED IN FORTRAN 60 OR 63

IY = 3125 * IY

IY = IY - (IY/67108864) * 67108864

FY = IY

KRAN = FY/67108864. * (B-A) + A

RETURN

END

FINIS

EXECUTER.

0378

0379

0380

0381

0382

0383

0384

0385

0386

0387

0388

0389

0390

0391

5	5						
1	1	QJ(1)	=2.64762775E-02	4	1	QJ(52)	=3.79764295E-02
1	2	QJ(2)	=3.02020831E-02	4	2	QJ(53)	=3.03224061E-02
1	3	QJ(3)	=1.72436823E-02	4	3	QJ(54)	=1.16230913E-02
1	4	QJ(4)	=6.68152510E-03	4	4	QJ(55)	=2.87808733E-03
1	5	QJ(5)	=2.02525670E-03	4	5	QJ(56)	=5.26888829E-04
1	6	QJ(6)	=5.29756823E-04	4	6	QJ(57)	=7.84720454E-05
1	7	QJ(7)	=1.30002033E-04	4	7	QJ(58)	=1.04516559E-05
1	8	QJ(8)	=3.22530060E-05	4	8	QJ(59)	=1.39145026E-06
1	9	QJ(9)	=8.58857000E-06	4	9	QJ(60)	=2.08960624E-07
1	10	QJ(10)	=2.86285667E-06	4	10	QJ(61)	=4.35334634E-08
1	11	QJ(11)	=8.34999861E-07	4	11	QJ(62)	=7.25557723E-09
1	12	QJ(12)	=2.08749965E-07	4	12	QJ(63)	=9.06947154E-10
1	13	QJ(13)	=4.34895761E-08	4	13	QJ(64)	=7.55789295E-11
1	14	QJ(14)	=7.24826268E-09	4	14	QJ(65)	=3.14912206E-12
1	15	QJ(15)	=9.06032836E-10	4	15	QJ(66)	= 0
1	16	QJ(16)	=7.55027363E-11	4	16	QJ(67)	= 0
1	17	QJ(17)	=3.14594735E-12	4	17	QJ(68)	= 0
2	1	QJ(18)	=3.02519433E-02	5	1	QJ(69)	=4.91642833E-01
2	2	QJ(19)	=3.07381552E-02	5	2	QJ(70)	=1.43718127E-01
2	3	QJ(20)	=1.54385405E-02	5	3	QJ(71)	=2.69531853E-02
2	4	QJ(21)	=5.18865107E-03	5	4	QJ(72)	=3.78709962E-03
2	5	QJ(22)	=1.34374437E-03	5	5	QJ(73)	=4.23732653E-04
2	6	QJ(23)	=2.96164302E-04	5	6	QJ(74)	=3.99381550E-05
2	7	QJ(24)	=6.07227021E-05	5	7	QJ(75)	=3.41914062E-06
2	8	QJ(25)	=1.26203649E-05	5	8	QJ(76)	=2.98561135E-07
2	9	QJ(26)	=2.86289503E-06	5	9	QJ(77)	=3.11088104E-08
2	10	QJ(27)	=8.35011051E-07	5	10	QJ(78)	=4.95114755E-09
2	11	QJ(28)	=2.08752763E-07	5	11	QJ(79)	=5.92108046E-10
2	12	QJ(29)	=4.34901589E-08	5	12	QJ(80)	=4.72886874E-11
2	13	QJ(30)	=7.24835981E-09	5	13	QJ(81)	=1.89139810E-12
2	14	QJ(31)	=9.06044977E-10	5	14	QJ(82)	= 0
2	15	QJ(32)	=7.55037481E-11	5	15	QJ(83)	= 0
2	16	QJ(33)	=3.14598950E-12	5	16	QJ(84)	= 0
2	17	QJ(34)	= 0	5	17	QJ(85)	= 0
3	1	QJ(35)	=3.40989043E-02	6	1	QJ(86)	= 0
3	2	QJ(36)	=3.07597531E-02	6	2	QJ(87)	= 0
3	3	QJ(37)	=1.35265107E-02	6	3	QJ(88)	= 0
3	4	QJ(38)	=3.91531781E-03	6	4	QJ(89)	= 0
3	5	QJ(39)	=8.57092801E-04	6	5	QJ(90)	= 0
3	6	QJ(40)	=1.56625190E-04	6	6	QJ(91)	= 0
3	7	QJ(41)	=2.62271034E-05	6	7	QJ(92)	= 0
3	8	QJ(42)	=4.44105168E-06	6	8	QJ(93)	= 0
3	9	QJ(43)	=8.35125955E-07	6	9	QJ(94)	= 0
3	10	QJ(44)	=2.08781489E-07	6	10	QJ(95)	= 0
3	11	QJ(45)	=4.34961435E-08	6	11	QJ(96)	= 0
3	12	QJ(46)	=7.24935725E-09	6	12	QJ(97)	= 0
3	13	QJ(47)	=9.06169656E-10	6	13	QJ(98)	= 0
3	14	QJ(48)	=7.55141380E-11	6	14	QJ(99)	= 0
3	15	QJ(49)	=3.14642242E-12	6	15	QJ(100)	= 0
3	16	QJ(50)	= 0	6	16	QJ(101)	= 0
3	17	QJ(51)	= 0	6	17	QJ(102)	= 0

7	1	QJ(103)=	0
7	2	QJ(104)=	0
7	3	QJ(105)=	0
7	4	QJ(106)=	0
7	5	QJ(107)=	0
7	6	QJ(108)=	0
7	7	QJ(109)=	0
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7	12	QJ(114)=	0
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EAN A/C FLYING =3.1666E 00

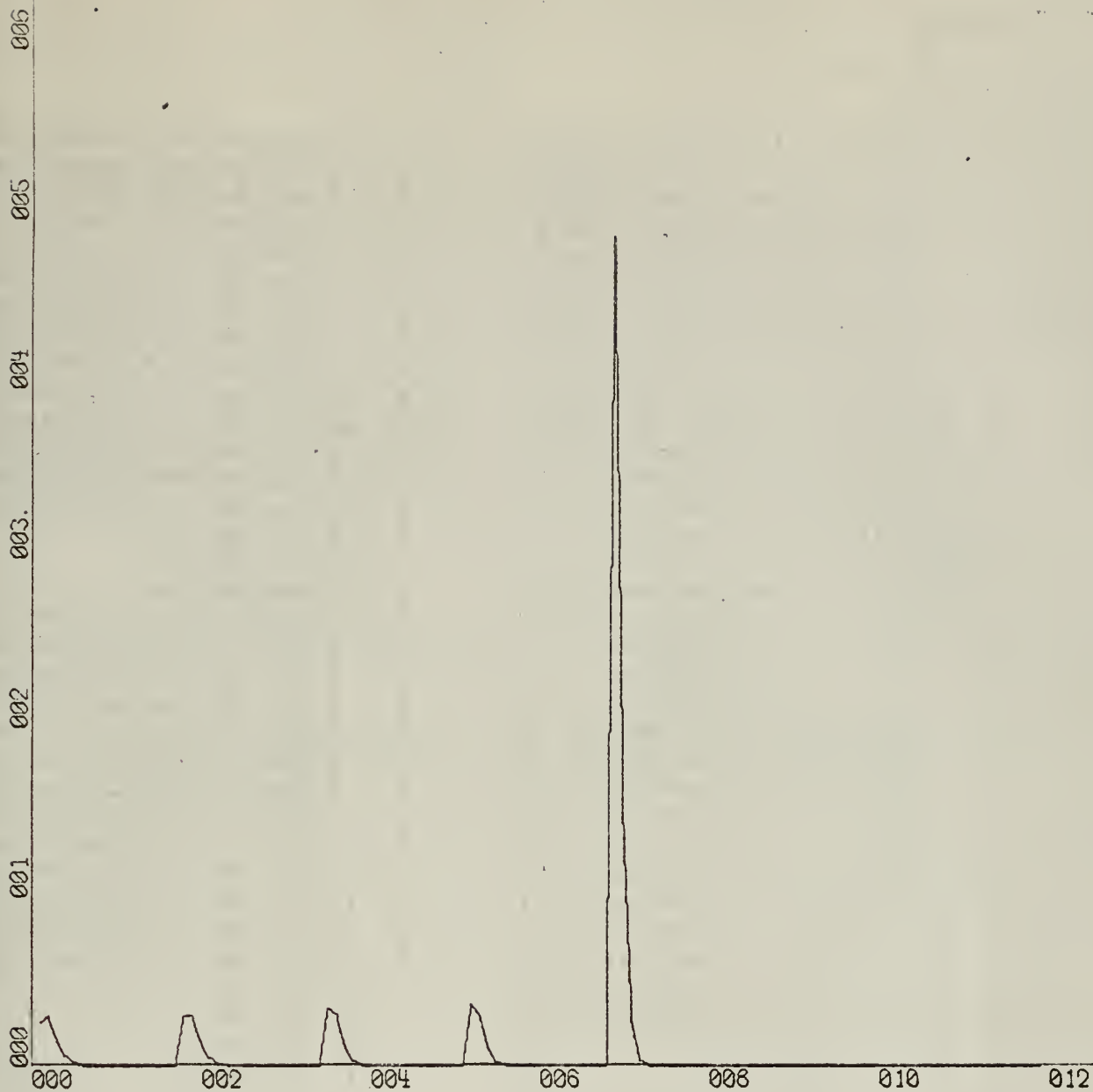
GRAPH TITLED

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HAS BEEN PLOTTED.

APPENDIX IV

SAMPLE RESULTS

The following pages present the values of the elements of the probability distribution vector (QJ) and its graphical plot for five consecutive iterations, i. e., $Q \times P^n$ for $n = 1, 2, \dots, 5$. The inputs are those shown on the first page of Appendix III between statement No. 30 and No. 31. The printouts of the transition matrices and their computational elements are omitted. The plot was made using the DRAW subroutine in the U. S. Naval Postgraduate School computer facility library. Each vector printout contains the values of all 119 states possible (7×17) and is headed by the past value of $A + 1$ and the next value of $A + 1$. The two indices preceding each element represent $\beta + 1$ and $j + 1$, in the notation of section 3. For example, in the first row on the next page, the "1 1" indicates that the probability of being in state (0, 0) after one iteration is $\hat{=} .026$, where the value of A is 4 over the first iteration. Each graph is labeled with the expected value of a/c flying, the number of maintenance spots available, the vector number (T), total number of a/c available (N), and the desired number of a/c on station (A). The "E" notation indicates the power of 10 to multiply by. This sample run demonstrates the loss in total a/c and variable a/c on station.



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Y-SCALE = 1.00E-01 UNITS/INCH.

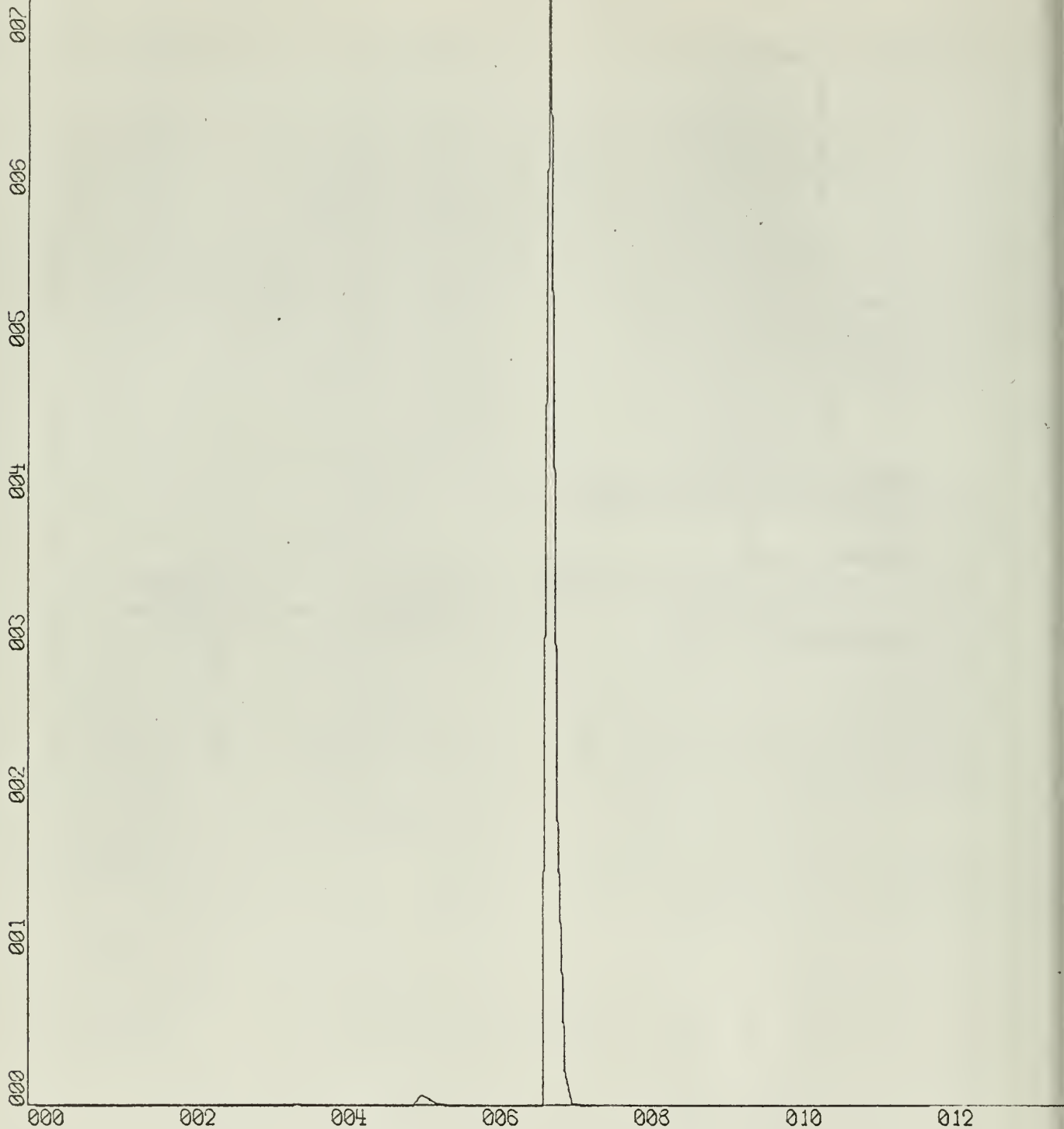
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 J VS QJ VECTOR N = 1.60E+01 A = 4.00E+00

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1	2	QJ(2)=7.94491028E-06	4	2	QJ(53)=4.45693238E-03	
1	3	QJ(3)=3.41380533E-06	4	3	QJ(54)=1.28968676E-03	
1	4	QJ(4)=9.83913729E-07	4	4	QJ(55)=2.35500532E-04	
1	5	QJ(5)=2.20530359E-07	4	5	QJ(56)=3.11156232E-05	
1	6	QJ(6)=4.28567319E-08	4	6	QJ(57)=3.30021596E-06	
1	7	QJ(7)=7.97994342E-09	4	7	QJ(58)=3.15123146E-07	
1	8	QJ(8)=1.56348438E-09	4	8	QJ(59)=3.13674353E-08	
1	9	QJ(9)=3.46529220E-10	4	9	QJ(60)=3.81044171E-09	
1	10	QJ(10)=1.01221242E-10	4	10	QJ(61)=6.98663591E-10	
1	11	QJ(11)=2.58173003E-11	4	11	QJ(62)=1.02568738E-10	
1	12	QJ(12)=5.63385201E-12	4	12	QJ(63)=1.13092255E-11	
1	13	QJ(13)=1.02295616E-12	4	13	QJ(64)=8.32962386E-13	
1	14	QJ(14)=1.48426307E-13	4	14	QJ(65)=3.07550593E-14	
1	15	QJ(15)=1.61419113E-14	4	15	QJ(66)=	0
1	16	QJ(16)=1.17030587E-15	4	16	QJ(67)=	0
1	17	QJ(17)=4.24536876E-17	4	17	QJ(68)=	0
2	1	QJ(18)=1.69105396E-04	5	1	QJ(69)=7.70076632E-01	
2	2	QJ(19)=1.31058642E-04	5	2	QJ(70)=1.88159529E-01	
2	3	QJ(20)=4.93425071E-05	5	3	QJ(71)=2.32480412E-02	
2	4	QJ(21)=1.22416819E-05	5	4	QJ(72)=1.93788734E-03	
2	5	QJ(22)=2.31695297E-06	5	5	QJ(73)=1.22762068E-04	
2	6	QJ(23)=3.73462197E-07	5	6	QJ(74)=6.39420794E-06	
2	7	QJ(24)=5.70661929E-08	5	7	QJ(75)=3.02738808E-07	
2	8	QJ(25)=9.22154646E-09	5	8	QJ(76)=1.56312395E-08	
2	9	QJ(26)=1.72702939E-09	5	9	QJ(77)=1.13519205E-09	
2	10	QJ(27)=4.41481881E-10	5	10	QJ(78)=1.50627061E-10	
2	11	QJ(28)=9.65869676E-11	5	11	QJ(79)=1.51200491E-11	
2	12	QJ(29)=1.75883018E-11	5	12	QJ(80)=1.02093434E-12	
2	13	QJ(30)=2.56023495E-12	5	13	QJ(81)=3.47880865E-14	
2	14	QJ(31)=2.79434001E-13	5	14	QJ(82)=	0
2	15	QJ(32)=2.03393035E-14	5	15	QJ(83)=	0
2	16	QJ(33)=7.41007441E-16	5	16	QJ(84)=	0
2	17	QJ(34)=	5	17	QJ(85)=	0
3	1	QJ(35)=1.42525319E-03	6	1	QJ(86)=	0
3	2	QJ(36)=9.83757697E-04	6	2	QJ(87)=	0
3	3	QJ(37)=3.24331350E-04	6	3	QJ(88)=	0
3	4	QJ(38)=6.90234010E-05	6	4	QJ(89)=	0
3	5	QJ(39)=1.09403657E-05	6	5	QJ(90)=	0
3	6	QJ(40)=1.44032446E-06	6	6	QJ(91)=	0
3	7	QJ(41)=1.76353392E-07	6	7	QJ(92)=	0
3	8	QJ(42)=2.28131763E-08	6	8	QJ(93)=	0
3	9	QJ(43)=3.51078290E-09	6	9	QJ(94)=	0
3	10	QJ(44)=7.70559333E-10	6	10	QJ(95)=	0
3	11	QJ(45)=1.40851171E-10	6	11	QJ(96)=	0
3	12	QJ(46)=2.05928494E-11	6	12	QJ(97)=	0
3	13	QJ(47)=2.25869291E-12	6	13	QJ(98)=	0
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3	15	QJ(49)=6.05799194E-15	6	15	QJ(100)=	0
3	16	QJ(50)=	6	16	QJ(101)=	0
3	17	QJ(51)=	6	17	QJ(102)=	0

7	1	QJ(103)=	0
7	2	QJ(104)=	0
7	3	QJ(105)=	0
7	4	QJ(106)=	0
7	5	QJ(107)=	0
7	6	QJ(108)=	0
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 HAS BEEN PLOTTED.



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Y-SCALE = 1.00E-01 UNITS/INCH.

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1	4	QJ(4)	=1.91420746E-06
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1	7	QJ(7)	=7.77939446E-09
1	8	QJ(8)	=1.18756605E-09
1	9	QJ(9)	=2.11283841E-10
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1	11	QJ(11)	=1.08528620E-11
1	12	QJ(12)	=1.87739309E-12
1	13	QJ(13)	=2.57762883E-13
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1	15	QJ(15)	=1.77499822E-15
1	16	QJ(16)	=5.92695364E-17
1	17	QJ(17)	= 0
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2	2	QJ(19)	=3.26173533E-04
2	3	QJ(20)	=1.04905695E-04
2	4	QJ(21)	=2.16142330E-05
2	5	QJ(22)	=3.28306882E-06
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2	7	QJ(24)	=4.68083562E-08
2	8	QJ(25)	=5.62586373E-09
2	9	QJ(26)	=8.10258169E-10
2	10	QJ(27)	=1.69317736E-10
2	11	QJ(28)	=2.92764116E-11
2	12	QJ(29)	=4.01871708E-12
2	13	QJ(30)	=4.10297009E-13
2	14	QJ(31)	=2.76746676E-14
2	15	QJ(32)	=9.24174352E-16
2	16	QJ(33)	= 0
2	17	QJ(34)	= 0
3	1	QJ(35)	=3.67589958E-03
3	2	QJ(36)	=2.22029313E-03
3	3	QJ(37)	=6.25702594E-04
3	4	QJ(38)	=1.10292585E-04
3	5	QJ(39)	=1.38900387E-05
3	6	QJ(40)	=1.38024736E-06
3	7	QJ(41)	=1.21057358E-07
3	8	QJ(42)	=1.09262020E-08
3	9	QJ(43)	=1.21362937E-09
3	10	QJ(44)	=2.09057619E-10
3	11	QJ(45)	=2.86055733E-11
3	12	QJ(46)	=2.91366712E-12
3	13	QJ(47)	=1.96282619E-13
3	14	QJ(48)	=6.55549374E-15
3	15	QJ(49)	= 0
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3	17	QJ(51)	= 0
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4	2	QJ(53)	=9.01023249E-03
4	3	QJ(54)	=2.23842146E-03
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4	5	QJ(56)	=3.57140995E-05
4	6	QJ(57)	=2.82593562E-06
4	7	QJ(58)	=1.85894467E-07
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4	9	QJ(60)	=9.51529557E-10
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4	11	QJ(62)	=1.28940846E-11
4	12	QJ(63)	=8.53320924E-13
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6	2	QJ(87)	=4.52895010E-02
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6	6	QJ(91)	=3.92418576E-06
6	7	QJ(92)	=1.47699319E-07
6	8	QJ(93)	=4.43630448E-09
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6	10	QJ(95)	=9.80396467E-12
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6	12	QJ(97)	= 0
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6	14	QJ(99)	= 0
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7 4 QJ(106)=1.37794578E-03
7 5 QJ(107)=6.83380035E-05
7 6 QJ(108)=2.32262953E-06
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7 8 QJ(110)=8.80653967E-10
7 9 QJ(111)=1.18536160E-11
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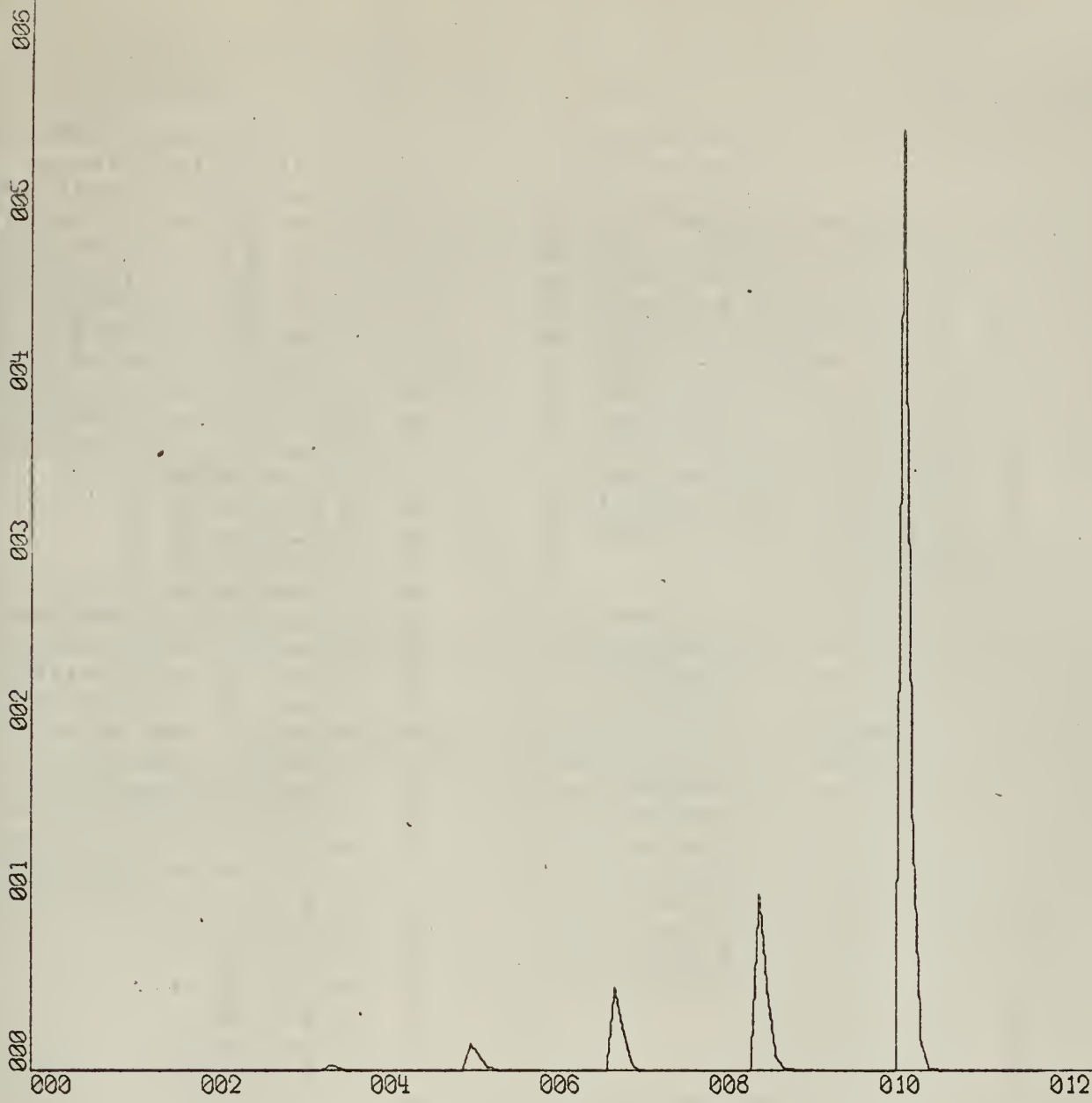
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HAS BEEN PLOTTED.



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T = 3.00E

J VS QJ VECTOR

N = 1.50E+01

A = 6.00

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1	4	QJ(4)	=1.37453763E-05
1	5	QJ(5)	=1.79617995E-06
1	6	QJ(6)	=1.89112432E-07
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1	12	QJ(12)	=9.07642678E-13
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1	14	QJ(14)	=5.27326607E-15
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1	17	QJ(17)	= 0
2	1	QJ(18)	=5.18126447E-03
2	2	QJ(19)	=2.86433504E-03
2	3	QJ(20)	=7.33497609E-04
2	4	QJ(21)	=1.16340651E-04
2	5	QJ(22)	=1.30125431E-05
2	6	QJ(23)	=1.13029797E-06
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2	8	QJ(25)	=6.65517506E-09
2	9	QJ(26)	=6.57969504E-10
2	10	QJ(27)	=1.05656138E-10
2	11	QJ(28)	=1.34342964E-11
2	12	QJ(29)	=1.26766173E-12
2	13	QJ(30)	=7.88829950E-14
2	14	QJ(31)	=2.42720088E-15
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2	17	QJ(34)	= 0
3	1	QJ(35)	=2.69210521E-02
3	2	QJ(36)	=1.34302284E-02
3	3	QJ(37)	=3.06604116E-03
3	4	QJ(38)	=4.26067113E-04
3	5	QJ(39)	=4.06841633E-05
3	6	QJ(40)	=2.90134684E-06
3	7	QJ(41)	=1.70443704E-07
3	8	QJ(42)	=9.76651595E-09
3	9	QJ(43)	=7.11954817E-10
3	10	QJ(44)	=9.16366982E-11
3	11	QJ(45)	=8.74738102E-12
3	12	QJ(46)	=5.50237415E-13
3	13	QJ(47)	=1.70994033E-14
3	14	QJ(48)	= 0
3	15	QJ(49)	= 0
3	16	QJ(50)	= 0
3	17	QJ(51)	= 0
4	1	QJ(52)	=8.03893389E-02
4	2	QJ(53)	=3.58842827E-02
4	3	QJ(54)	=7.24458728E-03
4	4	QJ(55)	=8.75756513E-04
4	5	QJ(56)	=7.09643835E-05
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4	8	QJ(59)	=7.45697706E-09
4	9	QJ(60)	=3.67973791E-10
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4	11	QJ(62)	=2.27683580E-12
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4	16	QJ(67)	= 0
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5	3	QJ(71)	=2.20943200E-02
5	4	QJ(72)	=1.73104557E-03
5	5	QJ(73)	=9.18785067E-05
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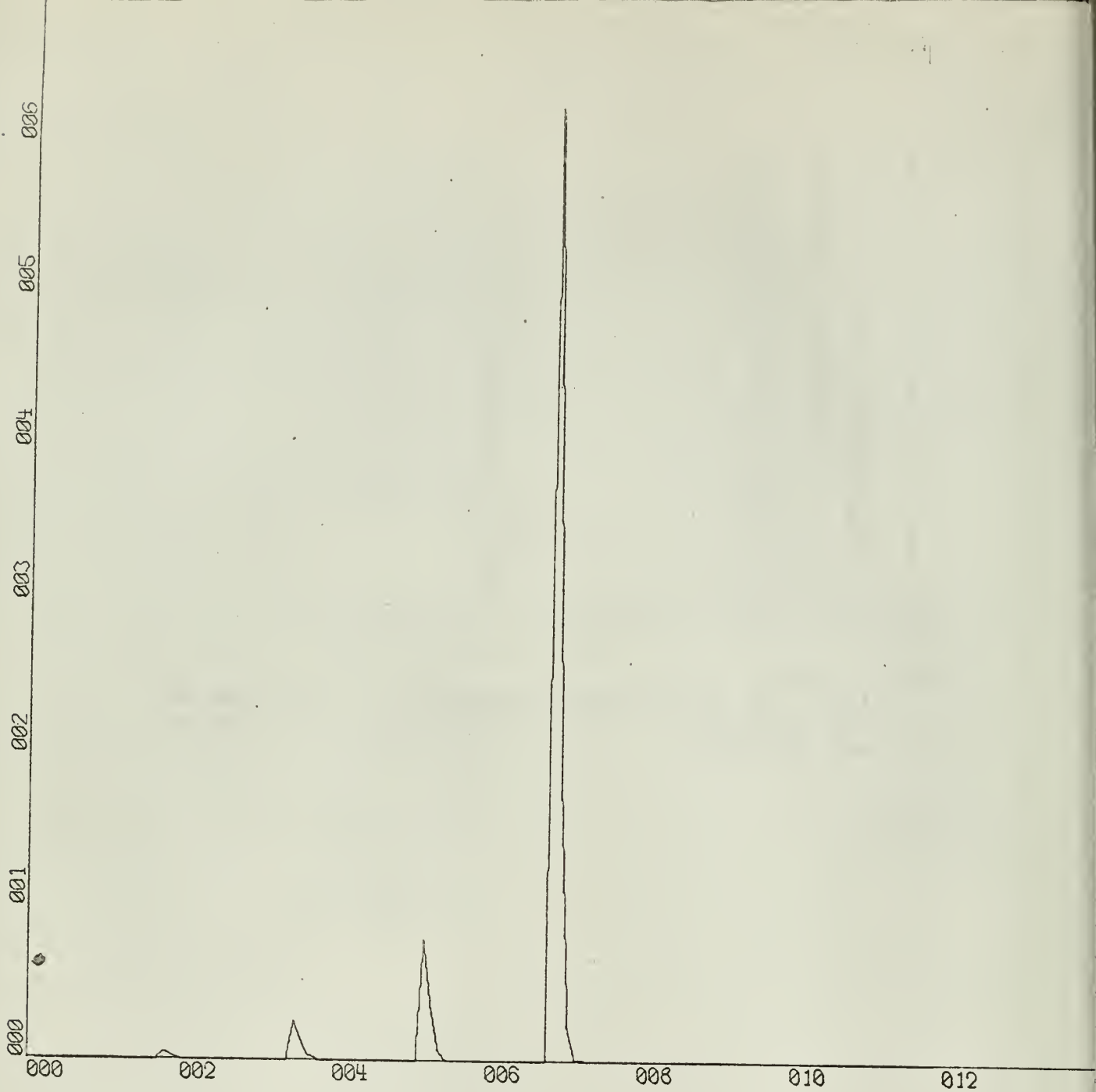
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J VS QJ VECTOR N = 1.40E+01 A = 4.00E+00

HAS BEEN PLOTTED.



X-SCALE = 2.00E+01 UNITS/INCH.

Y-SCALE = 1.00E-01 UNITS/INCH.

$E(A/C) = 3.76E+00$ SPOTS = $8.00E+00$

$T = 4.00E-$

J VS QJ VECTOR

$N = 1.40E+01$

$A = 4.00$

Q VECTOR CASE 5

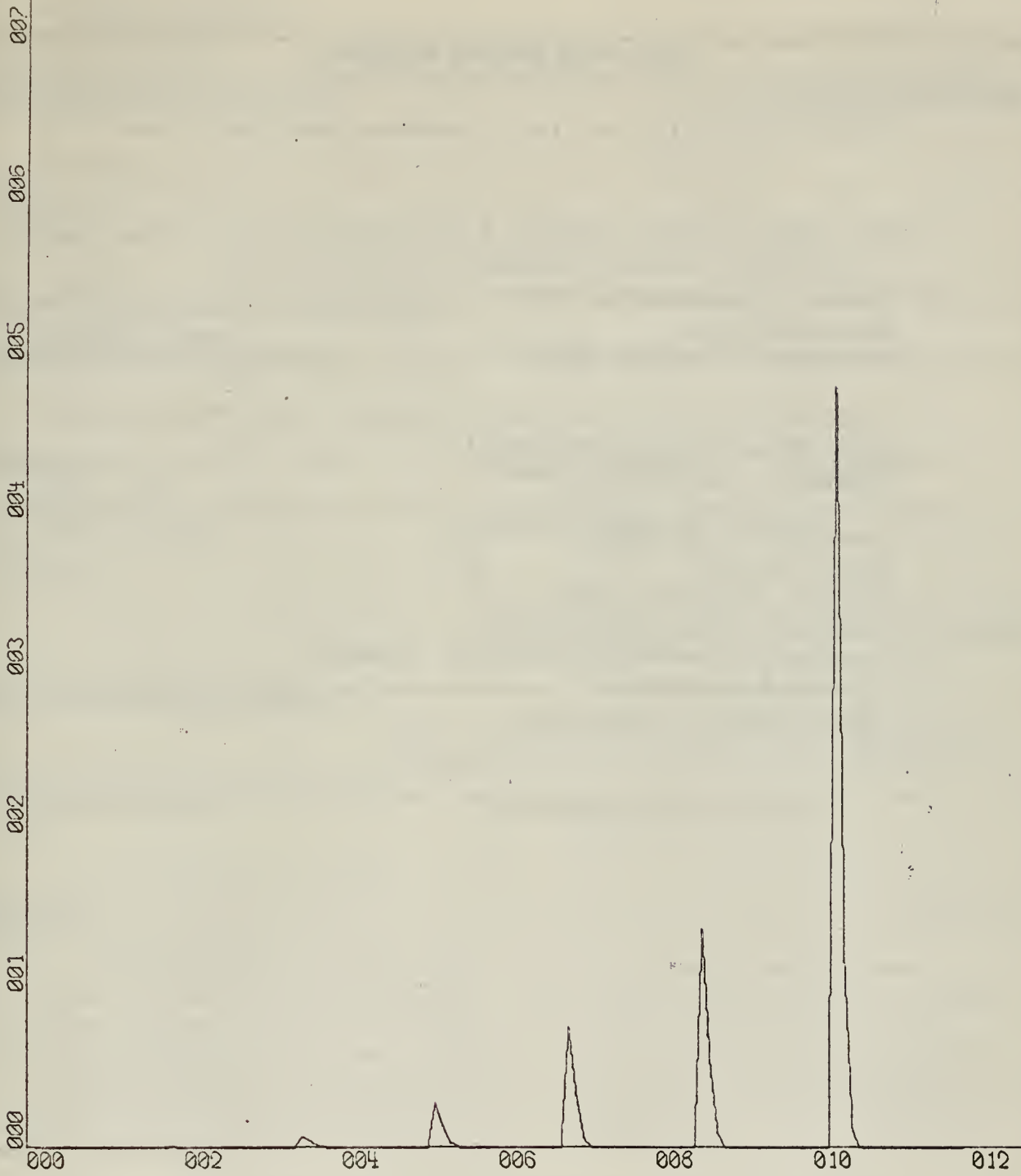
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1	11	QJ(11)	=4.80498022E-12
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		2b. GROUP	
3. REPORT TITLE AN ADAPTATION OF A MARKOV CHAIN MODEL FOR ANTISUBMARINE WARFARE CARRIER AIRCRAFT			
4. DESCRIPTIVE NOTES (Type of report and inclusive dates) Master's Thesis			
5. AUTHOR(S) (Last name, first name, initial) Lanman, George M., Lieutenant Commander, U. S. Navy			
6. REPORT DATE May, 1966		7a. TOTAL NO. OF PAGES 65	7b. NO. OF REFS 5
8a. CONTRACT OR GRANT NO.		8a. ORIGINATOR'S REPORT NUMBER(S)	
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13. ABSTRACT <p>It is the purpose of this paper to develop a useful mathematical model of ASW aircraft availability. The increasing emphasis of systems studies dictates the use of accurate and representative models of the ASW systems. At present, many studies are using essentially the same models developed during World War II. This paper is an attempt to make use of advanced theory in a more powerful and flexible model and to make the use of the model practical and verifiable.</p> <p>The writer adapted the time homogeneous bivariate model as developed by F. C. Collins. This is a discrete time Markov process with a stochastic matrix of transition probabilities wherein the maintenance process is modeled as a pulsed input multiple server queue.</p> <p>The model was programmed in FORTRAN 63 on the CDC 1604 and then modified to allow for variability in the input parameters. Other modifications include an increase in the size of the model to accommodate a 16-aircraft squadron, the largest ASW squadron at present, and an explicit form solution to the maintenance queueing equations.</p>			

14. KEY WORDS	LINK A		LINK B		LINK C	
	ROLE	WT	ROLE	WT	ROLE	WT
ASW Aircraft						
Computer Markov Model						
Availability						

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